Original Research

The Evaluation, Dynamic Evolutionary Characteristics and Influencing Factors of Green Innovation Efficiency in China

Minjie Li1 , Shuangjiao Lin² , Yihui Chen3*

¹School of Economics and Trade, Fujian Jiangxia University, Fuzhou 350108, China 2 Institute of Economics and Management, Xiamen University of Technology, Xiamen 361024, China 3 School of Digital Economy, Fujian Agriculture and Forestry University, Fuzhou 350002, China

> *Received: 18 February 2024 Accepted: 27 April 2024*

Abstract

Green innovation efficiency (GIE), which reflects the relationship between inputs and outputs that consider environmental impacts, is critical to China's realization of green and sustainable development. In view of the differences in economic development and resource endowment across provinces, it is necessary to measure and evaluate the GIE in a rational manner. In this paper, we study the evaluation, dynamic evolutionary characteristics, and influencing factors of GIE for 30 provinces in China over the period 2011-2022, using the super-SBM undesirable model, the spatial Markov chain model, and the geographically and temporally weighted regression model, respectively. Results show that the eastern region has a significantly higher GIE than the other regions, followed by the central region. Also, the state of provincial GIE in China is affected by the level of neighboring regions. GDP per capita, the marketization index, and industrial structure promoted GIE, while R&D expenditure, import and export trade, and the digital financial inclusion index showed negative results. There is also spatiotemporal heterogeneity in the effects of individual factors on GIE. Thus, a one-size-fits-all policy for GIE is highly unsuitable for the realities of individual provinces in China. We propose a number of targeted and differentiated policy recommendations that can be implemented.

Keywords: green innovation efficiency, super-SBM undesirable model, dynamic evolution, GTWR, China

Introduction

With regard to economic progress, social stability, and environmental protection, green is a common concern worldwide, as well as a way forward and a wise choice [1]. In the context of increasing global emissions of pollutants, particularly greenhouse gases, green innovation can help to promote economic development while maintaining environmental sustainability [2]. It is apparent that promoting green innovation has become a necessary route to achieving high-quality development [3]. Due to the high risks normally associated with technological innovation [4], green innovation cannot be achieved without the support of

^{*}e-mail: chenyihui@fafu.edu.cn

a large number of scientific and technical personnel and financial resources. According to the National Bureau of Statistics of China, the intramural expenditure on R&D and the full-time equivalent of R&D personnel in China reached 3.078 trillion yuan and 6.354 million man-years in 2022, respectively. As for industrial enterprises above the designated size, the percentage of enterprises with R&D activities in the total number of enterprises was as high as 37.3% in 2022, much higher than the 13.7% experienced in 2012. Evidently, both the government and enterprises have consistently invested heavily in the process of promoting green innovation. Against this backdrop, China has made many world-renowned achievements, such as supercomputing and manned spaceflight [5].

The high-risk attributes of green innovation signify that innovative activities are prone to failure and that technological transformation and upgrade of enterprises are costly. Therefore, high investments are not necessarily meant to imply high productivity. Accordingly, green innovation efficiency (GIE) has gradually become one of the crucial points of concern in both academia and practice. From the perspective of environmental concerns, GIE generally refers to an innovation efficiency that considers environmental benefits and economic outputs [6, 7]. Similarly, in terms of the input-output perspective, GIE refers to obtaining optimal innovation achievements with minimal resource and environmental costs [3, 8]. That is, a relatively low GIE may facilitate technological progress and upgrading within the short term, but it certainly hinders green technological innovation and environmental sustainability in the long term.

China's GIE is largely constrained by an excessive focus on economic benefits and a huge waste of resources in innovation activities. Additionally, the rigidity of technology innovation management regimes may also hinder the exploration of emerging and advanced technologies. In view of many external factors, such as the technology blockade from developed countries, China's GIE to date is still relatively low [9, 10]. Based on this, accurately evaluating the GIE of each province in China has become an important prerequisite for grasping the current state of green innovation. Commonly, the influencing factors of GIE are multiple and complex, involving various aspects of social prosperity and economic development. Meanwhile, inter-provincial differences due to resource endowments are also an issue that the Chinese government needs to take fully into account. This implies that there may be spatiotemporal heterogeneity in the impacts of a specific variable on GIE. It is apparently inappropriate to adopt a one-size-fits-all policy for different provinces. Thus, this study is also beneficial to the formulation of China's rational science and technology policies.

In summary, this paper addresses the following questions in light of the research gaps that exist in previous studies of China's GIE: What is the current level of provincial GIE in China, with full consideration of economic, social, and environmental aspects? What kind of spatial and temporal evolution has China's GIE experienced during the period of examination? Can the ranks of GIE in each province be transferred across ranks, and is this transfer related to the ranks of the surrounding provinces? For individual provinces, which factors have a positive and which have a negative impact on GIE, and does this vary from province to province?

Literature Review

In contrast to green innovation, which typically considers only the number of technological innovation outputs [11, 12], the evaluation of GIE is mainly concerned with the proportional relationship between green innovation inputs and outputs, i.e., efficiency [13, 14]. Among the input variables, labor and capital are the ones that have been fully considered by the most existing literature. For instance, Li [3], Song, and Han [15] employed human resources and financial investment as inputs. Besides, energy inputs are also significant input indicators considered by many scholars when measuring GIE. Proxies for energy inputs vary depending on the industry sector and the spatial scale of existing studies. Total energy consumption [8], total power supply [16], total electricity consumption [7, 17], and total gas supply [18] are the most commonly used proxy variables for energy inputs. On the methodological side, the data envelopment analysis (DEA) model is most popular when it comes to calculating relative efficiency [3]. However, traditional DEA models, such as the CCR-DEA model and the BCC-DEA model, have some shortcomings [19]. Specifically, these models ignore input and output slack variables and are unable to distinguish and compare the efficiency of the effective DMUs [3]. In this regard, Tone [20] and Tone [21] introduced the slacks-based measure (SBM) method and super-efficiency into traditional DEA models, respectively. Hence, the super efficiency SBM model has been widely applied to calculate GIE [6]. Additionally, as public concern regarding global environmental issues increases, the undesirable outputs are becoming a non-negligible part of the efficiency calculations. The outputs are categorized into desirable and undesirable outputs. In the calculation of GIE, desirable outputs consist mainly of economic benefits, and undesirable outputs mainly refer to environmental impacts and burdens caused by innovative activities. Undesirable outputs are commonly added to the calculation formula for GIE as part of the output [22]. However, there are still a few existing studies that calculate GIE by treating undesirable outputs as part of the inputs [23].

GIE involves input and output indicators that are multifaceted, and similarly, it is affected by many factors, including social, economic, political, geographic, and environmental [24]. For instance, theoretically, the urbanization level, which reflects the degree of social development, significantly promotes green technology

innovation activities [25]. Economic development is closely linked to infrastructure, construction, and technological research and development [26], which is conducive to attracting a large number of innovative entities [27]. Additionally, the government's governance, policy preferences, and decisions, such as policy support and fiscal decentralization, can also influence the effective improvement of GIE [7, 28]. Usually, local fiscal decentralization reflects the fiscal power of local governments, which in turn influences policy preferences. In economically developed regions, local fiscal decentralization may lead governments to pursue higher and faster short-term economic benefits at the expense of long-term environmental benefits, thereby suppressing GIE [28]. In addition, geographic factors are closely associated with a region's resource endowments, such as land and energy, as well as industrial structure and agglomeration, which clearly affect GIE. Geographic spatial linkages between regions have also been one of the hotspots for scholars in recent years [24]. Besides, environmental regulations, such as environmental pollution, control investment, and pollution removal rate, also significantly affect GIE [8, 16]. In terms of methodology for impact effects, the spatial econometric model [3], panel threshold model [16], quadratic assignment procedures [24], fuzzy set qualitative comparative analysis [29], geographical weighted regression model [30], generalized method of moment, and moderating effect model [31] have been applied to empirically examine the impact of influencing factors on GIE. However, many of the results obtained from the above methodologies fail to distinguish among the heterogeneity of impacts in both temporal and spatial terms, yielding one-size-fits-all policies. This shortcoming affects regional policy preferences, which in turn affects GIE.

There are three potential contributions to this study in comparison to existing studies. First, this study fully considers both the quantity and quality of green innovation as output variables for measuring GIE. Specifically, the number of patent applications is employed to indicate the quantity of green innovation, while the number of green invention patent grants is adopted to indicate the quality of green innovation. Second, the spatial Markov chain model is applied to dynamically analyze the spatial evolution of GIE, whereas most of the existing studies mainly selected a few specific years for static analysis. Third, the traditional methods employed in the majority of existing studies yield the same regression coefficients for each spatial location and time point. The geographically and temporally weighted regression (GTWR) model applied in this study can capture the parameter variability of different spatial units in both temporal and spatial dimensions.

Materials and Methods

Data Sources

Considering the availability of data, we selected 30 provinces in China (excluding Xizang, Hong Kong, Macao, and Taiwan) as the research objects in this paper. The time frame of this study is 2011-2022. Eventually, a total of 360 samples will be included in this study. The raw data for input and output variables employed to calculate GIE were obtained from the China Statistical Yearbook, the Chinese Research Data Services Platform, the China Statistical Yearbook on Science and Technology, and the China Statistical Yearbook on Environment. In addition, GDP per capita, industrial structure, and import and export trade could be obtained by calculating the raw data from the China Statistical Yearbook. The digital financial inclusion index was released by the Institute of Digital Finance at Peking University and the Research Institute at Ant Group. R&D expenditure was obtained from the Communiqué on National Expenditures on S&T released by the National Bureau of Statistics of China. Besides, the marketization index was produced by the China Market Index Database. The descriptive statistics of the selected influencing factors are shown in Table 1.

Definition and Measurement of Variables

Explained Variable

The super-SBM undesirable model is employed to evaluate China's GIE in this paper. The model involves 3 types of variables simultaneously: input variables, and desirable and undesirable output variables. Input

Table 1. The descriptive statistics of influencing factors.

Variables	Mean	St. D.	Min	Median	Max
GDPPC	5.872	3.066	1.602	5.064	19.031
DFI	243.928	107.640	18.330	255.931	460.691
RDE	1.743	1.164	0.410	1.455	6.830
MI	8.150	1.946	3.359	8.266	12.864
INS	1.266	0.722	0.518	1.112	5.310
IET	0.256	0.260	0.007	0.152	1.398

variables mainly include labor, capital, and energy. Hence, in this paper, input variables include R&D expenditure, R&D personnel, and energy consumption. We select intramural expenditure on R&D (10000 yuan) to represent R&D expenditure and choose the full-time equivalent of R&D personnel (man-year) to denote R&D personnel. Also, electricity consumption (100 million kWh) is selected as a proxy for energy consumption. In addition, desirable output variables consist of 3 parts: patent applications, green patent grants, and sales revenue of new products, considering both quantity and quality. We choose domestic patent applications (piece), green invention patent grants (piece), and sales revenue of new products of industrial enterprises above the designated size (10000 yuan) as proxies for desirable outputs, respectively. Furthermore, undesirable output variables are composed of 3 parts: wastewater discharge, waste gas emissions, and solid waste generated. In this paper, the chemical oxygen demand discharge in wastewater (10000 tons), sulfur dioxide emissions in waste gas (10000 tons), and common industrial solid wastes generated (10000 tons) are chosen to represent the undesirable output variables.

Explanatory Variables

As mentioned previously, GIE is influenced by many factors from various aspects. Hence, considering the availability and applicability of data, referring to [15, 19, 22, 23, 32], we choose economic development, marketization level, industrial structure, digital economy, government financial support, and trade openness as the influencing factors and explore the heterogeneity of their effects on GIE from temporal and spatial perspectives. Specifically, GDP per capita (GDPPC, 10000 yuan) is selected as the proxy of economic development; the digital financial inclusion index (DFI) is applied to represent the digital economy. Besides, the proportion of R&D expenditure to GDP (RDE) is used to denote government financial support, and the marketization index (MI) is chosen to represent the marketization level. Moreover, the ratio of the output value of the tertiary industry to the output value of the secondary industry (INS) is used to measure industrial structure, and the proportion of total import and export trade in goods to GDP (IET) is applied to measure trade openness.

Research Methods

The Super-SBM Undesirable Model

The super-SBM undesirable model has been demonstrated to be superior to the traditional DEA models in many research areas. It takes into account both undesirable outputs and allows for more in-depth comparisons of effective DMUs. Hence, it was adopted to evaluate China's GIE in this paper. Referring to [33], the super-SBM undesirable model can be constructed as follows:

$$
\min \rho = \frac{1 + \frac{1}{m} \sum_{m=1}^{M} \frac{s_m^x}{x_{jm}^t}}{1 - \frac{1}{l + h} \left(\sum_{l=1}^{L} \frac{s_l^y}{y_{jl}^t} + \sum_{h=1}^{H} \frac{s_h^b}{b_{jh}^t} \right)}
$$
\n
$$
s.t. \begin{cases} x_{jm}^t \ge \sum_{j=1, j \neq 0}^{n} \lambda_j^t x_{jm}^t + s_m^x \\ y_{jl}^t \ge \sum_{j=1, j \neq k}^{n} \lambda_j^t y_{jl}^t - s_l^y \\ b_{jh}^t \ge \sum_{j=1, j \neq k}^{n} \lambda_j^t b_{jh}^t + s_h^b \\ \lambda_j^t \ge 0, s_m^x \ge 0, s_l^y \ge 0, j = 1, \cdots, n \end{cases} (1)
$$

where *ρ* represents China's GIE. The parameter *λ* represents the constant vector. Besides, x_j^t , y_j^t , and b_j^t denotes inputs and desirable and undesirable outputs, respectively. In addition, *m*, *l*, and *h* represent the number of inputs, desirable outputs, and undesirable outputs, while s_m^x , s_l^y , and s_h^b are the corresponding slack vectors, respectively.

Spatial Markov Chain Model

When exploring the distributional evolution of variables, as an extension of the traditional Markov chain model, the spatial Markov chain model can demonstrate the dynamic evolutionary characteristics of variables and explore spatial spillover effects [34]. Currently, the spatial Markov chain model has been widely employed in various academic fields, such as air pollution [35], urban health development efficiency [36], and urban resilience [37]. Hence, a spatial Markov chain model is used in this paper to explore the dynamic evolutionary characteristics of GIE. Besides, before executing the spatial Markov chain model, we first divided GIE into 5 levels based on the natural breaks (Jenks) method, i.e., the levels were employed to represent GIE of different provinces in different years. In the traditional Markov chain model, p_{ij} is usually used to denote the probability that this stage is in rank *i*, and the next stage transitions to rank *j*. Therefore, referring to Du et al. [34], $p_{ii}(k)$ is used to represent the probability that this stage is in rank *i*, and the next stage transitions to rank *j* while the adjacent regions are in rank *k* at this stage. Moreover, the spatial weight matrix involved in the spatial Markov chain model is generated on the basis of the queen adjacency principle in this paper.

Geographically and Temporally Weighted Regression

Traditional econometric methods, such as geographically weighted regression and spatial econometric models, ignore the role of temporal or spatial factors in the analysis. In order to account for both temporal and spatial nonstationarity and to compensate for the shortcomings of traditional econometric methods, Huang et al. [38] proposed the GTWR model to analyze house prices. Currently, the GTWR model has been widely applied in the fields of environmental science [39, 40] and economic analysis [41, 42], etc., and has been demonstrated to be superior to other traditional econometric models. Thus, referring to Chen et al. [43], the GTWR model can be constructed as follows:

$$
GIE_{ii} = \beta_0(u_i, v_i, t_i) + \beta_1(u_i, v_i, t_i) \times GDPPC_{it} + \beta_2(u_i, v_i, t_i)
$$

$$
\times DFI_{it} + \beta_3(u_i, v_i, t_i) \times RDE_{it}
$$

$$
+ \beta_4(u_i, v_i, t_i) \times MI_{it} + \beta_5(u_i, v_i, t_i) \times INS_{it}
$$

$$
+ \beta_6(u_i, v_i, t_i) \times IET_{it} + \varepsilon_{it}
$$

(2)

where u_i , v_i , and t_i represent the longitude coordinate, latitude coordinate, and specific year, respectively; *βⁱ* denotes the constant and regression coefficients, while *ε*^{*it*}</sup> is the error term.

Results and Discussion

Evaluation of China's GIE

Evaluation Results of GIE

The super-SBM undesirable model is employed to evaluate each province's GIE in China during 2011-2022 in this paper. The evaluation results of GIE in China by province are shown in Table 2. The lowest GIE was 0.137 in Inner Mongolia in 2011, while the highest GIE was 1.269 in Hainan in 2016. The latter is 9.263 times higher than the former, showing a very significant difference in GIE among provinces. In addition, GIE in most of the provinces does not show a consistent and clear trend, which fully illustrates the complexity of the evolution of GIE. This finding is consistent with Huang et al.'s [10] study that GIE in Hainan is almost always above 1.000, while the vast majority of China's provinces have GIEs below 1.000, i.e., a state of inefficiency, and GIEs in less developed regions are relatively low.

Spatiotemporal Evolution of GIE

The temporal evolution of GIE in China by region is shown in Fig. 1. Ordinarily, China's economic zones, which contain 31 provinces (excluding Hongkong, Macao, and Taiwan), can be divided into 4 regions: the eastern, central, western, and northeastern regions. The trends in the temporal evolution of the nationwide GIE and GIE in the individual regions remain highly consistent. For each region, the year 2016 is a critical juncture for significant changes in GIE. The nationwide GIE exhibits 3 distinct phases in the time dimension: 2011-2014, 2014-2016, and 2016-2022. GIE at these 3 phases successively shows a steady, sharp rise and fluctuating upward trend, respectively. In the long run, such a trend reflects the continued improvement of green innovation in China. This is consistent with the conclusions of Zhang et al. [25], which concluded that China's GIE has generally shown an increasing trend through time. The temporal evolution of GIE in the western region is strongly characterized by similar characteristics as the nationwide GIE, but the former always remains at a relatively low level after 2019 and does not exceed 0.5 until 2022.

In the eastern region, GIE is consistently and significantly higher than GIE in other regions, except in 2011. This finding confirms the conclusions of Zhao et al.'s [14] study that the eastern region has the highest average GIE. This is due to the fact that the eastern region has the best economic base and the highest level of technology, as well as being a magnet for highlevel talents from various regions. After 2016, GIE in the eastern region has been slowly and intermittently rising on the whole, and its dominance has become more pronounced, i.e. the phenomenon of the Matthew Effect has emerged. Additionally, the temporal evolution of GIE in the central region can be divided into 3 stages: 2011-2016, 2016-2019, and 2019-2022, with an accompanying rising-falling-rising characteristic. The economic base of the central region is second only to that of the eastern region, and it is also attractive to highlevel technical talents. Simultaneously, the central region is endowed with abundant resources, such as minerals, and GIE may be further substantially enhanced in the process of industrial green transformation. Fig. 1 also shows the significant potential for the central region to increase its GIE beyond 2019, which is different from the slow rise in the western region. This finding is in contrast to that of Zhao et al. [14] who concluded that the difference in GIE between the central and western regions is decreasing. One possible explanation is that the latter only calculates GIE up to 2020, and the regional division criteria for China are not consistent with this paper. The temporal evolution characteristics of GIE in the northeastern region are more complex compared to other regions. GIE underwent a downwardupward trend with a clear U-shaped pattern before 2016. By contrast, after 2016, GIE in the northeastern region has trended downward with significant fluctuations. The mass outflow of advanced technicians from the northeastern region and the large number of highly polluting heavy industries have inhibited local green innovation.

From a spatial perspective, Hainan has the highest average GIE of 1.056, the only province in the country with a GIE that exceeds 1. Recently, Hainan has been capitalizing on its natural tropical monsoon oceanic climate and island advantages to develop tourism resources, reduce heavily polluting heavy industries, and enhance the advantages of technological innovation in terms of greenness and sustainability. As indicated in Table 2, the average GIE is followed by Tianjin,

Province	2011	2012	2013	2014	2015	2016	2017	2018	2019	2020	2021	2022	Average
Hainan	1.245	1.007	1.089	0.815	0.784	1.269	1.063	1.080	1.096	1.093	1.088	1.046	1.056
Tianjin	0.627	0.671	1.011	0.861	1.002	1.080	0.816	1.005	1.039	1.033	0.767	1.008	0.910
Zhejiang	0.524	0.642	1.005	0.658	1.001	1.038	0.827	1.055	1.039	1.019	1.023	1.035	0.906
Shanghai	0.642	0.621	0.607	0.636	0.615	1.007	1.009	1.022	1.048	1.029	1.031	0.883	0.846
Beijing	0.494	0.508	0.554	0.597	0.643	1.047	1.021	1.028	1.001	1.013	1.086	1.121	0.843
Jiangsu	0.579	1.012	0.825	0.598	0.616	0.727	0.683	1.010	0.666	1.000	0.792	1.037	0.796
Qinghai	0.261	0.271	0.277	0.204	1.082	1.124	0.662	1.222	1.036	1.075	1.003	1.053	0.772
Guangdong	0.425	0.426	0.457	0.462	0.547	0.754	1.014	1.030	1.028	1.023	1.028	1.042	0.770
Chongqing	1.024	0.633	0.634	0.700	1.059	1.051	0.797	0.718	0.630	0.616	0.645	0.615	0.760
Anhui	0.434	0.465	0.491	0.494	0.595	1.015	0.847	1.054	0.617	0.706	0.701	1.002	0.702
Jilin	1.024	0.467	0.285	0.385	0.434	1.013	1.023	0.610	1.003	0.590	0.688	0.682	0.684
Hunan	0.434	0.477	0.533	0.542	0.627	1.003	0.721	0.580	0.523	0.523	0.592	0.597	0.596
Guangxi	0.311	0.335	0.432	0.457	0.632	1.105	1.043	0.683	0.446	0.552	0.553	0.511	0.588
Jiangxi	0.278	0.346	0.372	0.381	0.407	0.590	0.572	0.557	0.489	0.560	0.780	1.026	0.530
Shandong	0.387	0.412	0.431	0.417	0.409	0.489	0.486	0.471	0.475	0.504	0.556	1.018	0.505
Hubei	0.334	0.360	0.382	0.404	0.447	0.530	0.558	0.583	0.545	0.553	0.605	0.636	0.495
Guizhou	0.398	0.359	0.384	0.438	0.398	0.545	0.506	1.002	0.445	0.415	0.376	0.393	0.472
Ningxia	0.288	0.347	0.440	0.398	0.455	0.522	0.553	0.592	0.383	0.426	0.465	0.590	0.455
Fujian	0.372	0.366	0.357	0.332	0.415	0.579	0.549	0.624	0.465	0.500	0.442	0.438	0.453
Sichuan	0.312	0.342	0.376	0.379	0.452	0.574	0.592	0.504	0.407	0.416	0.417	0.415	0.432
Heilongjiang	0.341	0.367	0.388	0.340	0.366	0.421	0.474	0.558	0.457	0.461	0.517	0.494	0.432
Shaanxi	0.305	0.329	0.362	0.352	0.393	0.416	0.498	0.460	0.431	0.433	0.457	0.442	0.407
Liaoning	0.314	0.333	0.363	0.337	0.378	0.455	0.444	0.490	0.410	0.388	0.432	0.429	0.398
Gansu	0.315	0.322	0.336	0.339	0.329	0.397	0.441	0.475	0.449	0.417	0.379	0.408	0.384
Henan	0.241	0.237	0.329	0.334	0.375	0.409	0.461	0.523	0.360	0.432	0.424	0.421	0.379
Xinjiang	0.299	0.303	0.309	0.350	0.354	0.393	0.401	0.401	0.444	0.458	0.370	0.401	0.374
Hebei	0.222	0.255	0.270	0.268	0.294	0.347	0.361	0.421	0.437	0.450	0.571	0.468	0.364
Shanxi	0.224	0.228	0.236	0.226	0.260	0.372	0.374	0.421	0.378	0.384	0.421	0.457	0.332
Yunnan	0.267	0.292	0.304	0.309	0.281	0.370	0.355	0.423	0.312	0.315	0.357	0.364	0.329
Inner Mongolia	0.137	0.154	0.148	0.147	0.164	0.218	0.287	0.354	0.332	0.285	0.386	0.518	0.261
Average	0.435	0.430	0.466	0.439	0.527	0.695	0.648	0.699	0.613	0.622	0.632	0.685	0.574

Table 2. The evaluation results of GIE in China by province.

Zhejiang, Shanghai, and Beijing, with the former two exceeding 0.9 and the latter two exceeding 0.8. Of the 10 provinces in the eastern region, all 7 except Shandong, Fujian, and Hebei are ranked in the top 10 of GIE, which further confirms that GIE in the eastern region is much higher than in other regions as depicted in Fig. 1. Among the 10 provinces at the bottom of the rankings, Hebei belongs to the eastern region, Henan and Shanxi belong to the central region, Heilongjiang and Liaoning are from the northeastern region, while the other provinces are from the western region. The GIEs of these 10 provinces are below 0.45. The average GIE of Inner Mongolia is the lowest nationally, reaching only 0.261, accounting for only 24.716% of Hainan's GIE. While Inner Mongolia's average GIE is consistently the lowest nationwide until 2019, it is rising significantly, even surpassing 13 provinces to reach 0.518 in 2022.

Fig. 1. The temporal evolution of GIE in China by region.

Dynamic Evolutionary Characteristics of GIE

Traditional Markov Chain Model

The transition probability matrix of GIE based on the traditional Markov chain model is shown in Table 3. With the exception of those located at level 4, the GIEs of the provinces located at other levels in this stage remain stable, i.e., the probability of not experiencing a transition in the next stage exceeds 0.5. Provinces located at level 4 in this stage would jump to level 5 in the next stage with a probability of 0.591. Besides, there is a clear trend of upward transfer of level in the majority of provinces, but also a risk of level decline in some individual provinces, especially those whose GIE is at a high level at this stage. Overall, there is much more upward level transition than a downward level. Also, some of the transition probabilities away from the diagonal are not equal to zero, implying that there is a transition of GIE across levels in some provinces. This clearly provides a factual basis for those provinces that have taken appropriate measures with the expectation that GIE will be upward transfer across levels in the short term. Similar to this finding, the conclusions of the study by Xin et al. [44] confirm that green innovations across provinces stabilize to maintain the initial state. The difference, however, is that the latter confirms that there is no overstepping transfer in the level of green innovation in the provinces. This may be due to the fact that only four levels of provincial green innovation were categorized in Xin et al.'s [44] study.

Spatial Markov Chain Model

In comparison to the traditional Markov chain model, the spatial Markov chain model fully takes into account the spatial spillover effect due to geographic proximity and spatial lag conditions. In Table 4, the transition probability matrix of GIE is generated based on a spatial Markov chain model. Under different spatial lag conditions, the spatial Markov transition probability of GIE is different from the transition probability in Table 3, indicating that the transition probability of the level of GIE in the local region to the next stage is affected by the level of neighboring regions. Similarly, there are differences in the spatial Markov transition probabilities for GIEs at the same level under different spatial lag conditions. For instance, the probability that the local region has a GIE rating of level 2 and moves to level 3 in the next stage is 0.083, 0.146, 0.239, 0.077, and 0.000 for different spatial lag conditions, respectively. Additionally, in contrast to Table 3, there is still a high probability that a region's GIE will remain at its original level or jump upwards after accounting for spatial factors. For instance, when GIE of the local region and neighboring regions are at the same level in the current stage, the probability that GIE of the local region would remain at the same level in the next stage is 0.742, 0.854, 0.621, 0.100, and 0.778, respectively. Also, the level of GIE in the local region is influenced by neighboring regions, and there are relatively few circumstances in which there is a jump across levels and a transition downward. In some regions, the probability of a downward shift in the level of the local region's GIE is relatively high when adjacent to a province with a low level; conversely, the probability of an upward shift in the local region is also relatively high when adjacent to a province with a high level. The fact that the probability p_{23} (0.167) is less than $p_{21}(1)$ (0.250) and that $p_{23}(3)$ (0.239) is greater than p_{23} (0.167) proves this conclusion. To summarize, these findings confirm Xin et al.'s [44] conclusions that the state of China's provincial green innovation is significantly influenced by geospatial patterns.

p_{ij}	\boldsymbol{n}				4	
	62	0.645	0.339	0.000	0.000	0.016
	120	0.050	0.775	0.167	0.000	0.008
	65	0.000	0.123	0.569	0.108	0.200
	22	0.000	0.046	0.136	0.227	0.591
	61	0.000	0.033	0.115	0.197	0.655

Table 3. Traditional Markov transition probability matrix of GIE.

Table 4. Spatial Markov transition probability matrix of GIE.

\boldsymbol{k}	$p_{ij}(k)$	\boldsymbol{n}	$\mathbf{1}$	$\overline{2}$	\mathfrak{Z}	$\overline{4}$	$\sqrt{5}$
$\,1$	$\,1$	31	0.742	0.258	0.000	0.000	0.000
	$\sqrt{2}$	12	0.250	0.584	0.083	0.000	0.083
	$\overline{3}$	$\,1\,$	0.000	0.000	1.000	0.000	0.000
	$\overline{4}$	$\boldsymbol{0}$	0.000	0.000	0.000	0.000	0.000
	5	$\mathbf{1}$	0.000	$1.000\,$	0.000	0.000	0.000
$\sqrt{2}$	$\,1$	$27\,$	0.630	0.333	0.000	0.000	0.037
	$\sqrt{2}$	48	0.000	0.854	0.146	0.000	0.000
	3	19	0.000	0.105	0.474	0.158	0.263
	$\overline{4}$	5	0.000	0.000	0.000	0.400	0.600
	5	16	0.000	0.000	0.250	0.125	0.625
\mathfrak{Z}	$1\,$	$\overline{4}$	0.000	1.000	0.000	0.000	0.000
	$\overline{2}$	46	0.044	0.717	0.239	0.000	0.000
	\mathfrak{Z}	29	0.000	0.103	0.621	0.103	0.173
	$\overline{4}$	5	0.000	0.000	0.400	0.200	0.400
	5	13	0.000	0.077	0.077	0.308	0.538
$\overline{4}$	$\mathbf{1}$	$\boldsymbol{0}$	0.000	0.000	$0.000\,$	0.000	0.000
	$\sqrt{2}$	13	0.077	0.846	0.077	0.000	0.000
	\mathfrak{Z}	15	0.000	0.200	0.533	0.067	0.200
	$\overline{4}$	$10\,$	0.000	0.100	0.100	0.100	0.700
	$\mathfrak s$	$22\,$	0.000	0.000	0.046	0.227	0.727
$\sqrt{5}$	$\,1$	$\boldsymbol{0}$	0.000	0.000	0.000	0.000	0.000
	\overline{c}	$\mathbf{1}$	0.000	$1.000\,$	0.000	0.000	0.000
	$\overline{3}$	$\,1$	0.000	0.000	1.000	0.000	0.000
	$\overline{4}$	$\sqrt{2}$	0.000	0.000	0.000	0.500	0.500
	5	$\overline{9}$	0.000	$0.000\,$	0.111	0.111	0.778

Influences of Factors on China's GIE

Overall Characteristics of Influences

The GTWR model is adopted to explore the spatiotemporal heterogeneity of influences from factors on China's GIE in this paper. The GTWR model has

the values of adjusted R^2 , AICc, and residual sum of squares of 0.817, -312.122, and 4.643, respectively, both of which are significantly superior to the global OLS regression, geographically weighted regression model, and temporally weighted regression model. In regression results on the basis of the GTWR model, there is a separate regression coefficient for each variable for

Variables	Mean	Min	Q_{1}	Median	Q_3	Max
GDPPC	0.055	-0.106	-0.021	0.021	0.089	0.863
DFI	-0.000	-0.021	-0.001	-0.000	0.000	0.003
RDE	-0.032	-0.965	-0.147	-0.001	0.095	1.095
MI	0.051	-0.176	0.006	0.058	0.104	0.282
INS	0.094	-0.278	-0.086	0.048	0.267	0.859
IET	-0.034	-4.622	-0.224	-0.003	0.325	0.846

Table 5. The descriptive statistics of regression coefficients based on GTWR model.

each province at each time point. Table 5 depicts the descriptive statistics of regression coefficients based on the GTWR model. As shown in Table 5, in general, GDPPC, MI, and INS promoted GIE, while DFI, RDE, and IET inhibited GIE. In Table 5, the average of the regression coefficients of RDE on China's GIE is -0.032, indicating a negative effect of the former on the latter. One possible explanation is the relatively great differences in economic development, infrastructure, and resource endowment among provinces, leading to differences in the impact of R&D expenditure on GIE. There are 8 provinces where RDE consistently shows a positive effect on GIE, which are Henan, Anhui, Jiangsu, Hebei, Shanxi, Shanghai, Zhejiang, and Shandong. In comparison, in other provinces, the impact of RDE on GIE is sometimes negative, and in 8 of them, it is consistently negative. Geographically, the 8 provinces in which RDE exhibits a positive effect are spatially closely connected, showing obvious clustering characteristics. The 8 provinces that are consistently negative are mainly located in the less economically developed regions, i.e., the northeastern and western regions of China. These findings are different from Li et al. [22], possibly because we embed both temporal and spatial factors in the regression model, which is clearly a more scientific and rational treatment. Indeed, there are also conclusions from many studies, such as Song et al. [19], that support the findings of this paper that there are significant regional differences in the impact of influencing factors on China's provincial GIE. Therefore, an in-depth spatiotemporal heterogeneity analysis of impacts is necessary.

Temporal Heterogeneity of Influences

Fig. 2 displays the temporal heterogeneity of factors' influences on China's GIE. With the exception of 2021, the positive influence of GDPPC on GIE diminishes over time, from 0.087 in 2011 to 0.017 in 2022. Similarly, INS has consistently demonstrated a positive influence in any given year and has shown a V-shaped trend, with 2020 being the turning point. Moreover, with the exception of 2019, the impact of MI on GIE has also been consistently positive and shown a V-shaped trend. These results indicate that, on the one hand, the enhancement of the GIE is influenced by more and more factors, so that the influence of some traditional factors is gradually weakening. On the other hand, due to the impact of COVID-19 epidemiology, the marketization level and the tertiary sector are becoming increasingly critical to GIE.

Besides, the direction of DFI's effect on GIE appears to alternate negatively and positively. However, the negative effect of DFI on GIE tends to be close to zero, implying a very weak causal relationship between DFI and GIE. This is different from the findings of Li et al. [22] for the reason that Li et al. [22] focused on analyzing the influence of the digital economy on the industrial GIE, whereas this study does not differentiate between industries. According to Liang et al. [24], the digital economy contributes to sustainable innovation, but the adequacy of the digital economy infrastructure

Fig. 2. The temporal heterogeneity of influences.

significantly affects the full development of digital green innovation. Obviously, in China, the digital economy serves more as a crucial infrastructure to improve the efficiency of social operations and production, while its impact on advanced green technologies is minimal. This implies that China's provinces need to continue to promote the digital economy in order to effectively realize its contribution to GIE in the future. In addition, in Table 5, while both RDE and IET show an overall negative impact on GIE, Fig. 2 displays that the impacts of both RDE and IET on GIE shift from negative to positive, with 2019 as the turning point for the former and 2018 for the latter. China's continued investment in R&D in line with the current situation and the expansion of international exchanges and cooperation are effective ways to enhance GIE.

Fig. 3. The spatial heterogeneity of influences.

Spatial Heterogeneity of Influences

The spatial heterogeneity of factors' influences on China's GIE is illustrated in Fig. 3. The province with the largest positive effect of GDPPC on GIE is Heilongjiang, followed by Jilin and Liaoning, all located in the northeastern region. Other provinces with positive effects are mainly found in the west. In the eastern and central regions, where the economic development level is relatively high, GDPPC has a mainly negative effect on GIE. Indeed, the excessive pursuit of shortterm economic benefits by local governments leads to development at the expense of environmental pollution and ecological damage, which ultimately inhibits GIE. Economically underdeveloped regions need to continue to develop economies, while other regions need to curb the excessive pursuit of economic benefits. In terms of spatial distribution, the direction of DFI's impact on GIE is almost exactly opposite to that of GDPPC. The positive impact of DFI on GIE is mainly in the eastern and central regions, while the negative impact is mainly in the western and northeastern regions. Therefore, the western and northeastern regions need to strengthen and improve the infrastructure of the digital economy, so as to realize the role of the digital economy in the enhancement of GIE.

Similarly, the effect of RDE on GIE shows significant spatial heterogeneity. In the western and northeastern regions, RDE plays a predominantly negative role, while in other regions it plays a positive role. This suggests that the western and northeastern regions should improve the system of science and technology development and establish a good incentive mechanism to attract a large number of technical talents, high-tech enterprises, institutions of higher education, and other technological innovation subjects. As for other regions, it is imperative to continue to invest in R&D expenditure in order to sustain the important role of R&D in green technology innovation. As depicted in Fig. 3, MI positively contributes to GIE in all provinces except Ningxia, Gansu, and Sichuan. This provides a practical basis for local governments to enhance the development of product and factor markets and strengthen their linkages with markets.

The provinces with large positive contributions of INS to GIE are Guangdong, Hainan, and Guangxi, while the provinces with large negative contributions are Shanghai, Jiangsu, and Shandong. Most of the provinces where INS shows a negative impact on GIE are located in the eastern, central, and northeastern regions, while the western region mainly shows positive impacts. Hence, the northeastern region should take into account its natural environment and resource endowment and actively develop tourism resources, such as winter tourism programs, to reduce the longterm overdependence on the secondary industry. The eastern and central regions should also appropriately reduce highly polluting industries and improve the greenness and sustainability of technological innovation.

In addition, the negative impact of IET on GIE is mainly found in Qinghai, the northeastern region, and Xinjiang, all of which are relatively remote regions that are weak in terms of import and export trade. Therefore, these provinces can take advantage of national-level development strategies, such as the Belt and Road Initiative, to strengthen international trade exchanges and cooperation.

Conclusions and Policy Recommendations

In this paper, we employed the super-SBM undesirable model to evaluate the GIE of 30 provinces in China from 2011 to 2022 and explored the dynamic evolutionary characteristics of GIE and the spatiotemporal heterogeneity of the influencing factors using the spatial Markov chain model and the GTWR model, respectively. The conclusions we draw are as follows: First, the eastern region has a critically higher GIE than the other regions, followed by the central region, while the western and northeastern regions are not noticeably different in GIE with alternating phenomena. There are also significant differences in the average GIE between provinces due to differences in infrastructure and resource endowments. In the long run, the GIE of each region has been rising in fluctuation. Second, the Markov transition probability of GIE considering spatial spillover effects and spatial lag conditions differs from traditional transition probability, indicating that the transition probability of GIE is affected by the level of neighboring regions. Particularly in some regions, the probability of GIE's level decreasing is relatively high when adjacent to a province with a low level, and the probability of GIE's level moving upward is also relatively high when adjacent to a province with a high level. Third, as a whole, GDPPC, MI, and INS promoted GIE, while DFI, RDE, and IET inhibited GIE. Simultaneously, there is significant spatiotemporal heterogeneity in the effects of individual factors on GIE. GDPPC's positive impact on GIE diminishes over time, and the influences of both INS and MI on GIE show a V-shape trend in the time dimension. Besides, there is a very weak causal relationship between DFI and GIE, while the impacts of both RDE and IET on GIE shift from negative to positive.

In view of this, it is clear that a one-size-fits-all policy for GIE is highly unsuitable for the realities of individual provinces in China. Hence, based on the evolutionary dynamics of GIE in each province and the factors affecting it, we propose a number of targeted and differentiated policy recommendations that can be implemented. First, the governments have to take various measures to enhance the GIE of each province, which reflects the efficiency relationship between inputs and outputs. On the basis of ensuring the greenness and sustainability of technological innovation, it is necessary and feasible to appropriately reduce inputs or expand outputs. Thus, the governments can enhance GIE by

attracting various types of innovative entities, sounding and improving the technological innovation system, concerning themselves with geospatial linkages with other provinces, and maintaining a good relationship with the market. Second, individual provinces, particularly those in the eastern and central regions, should not pursue short-term economic benefits in promoting economic development, but balance the relationship between economic development and environmental protection. These provinces can rely on their good economic foundation to realize the transformation and upgrading of their industrial structure and vigorously develop tertiary and high-tech industries. Third, the provinces in the western and northeastern regions need to facilitate the infrastructure construction of the digital economy and increase funding for R&D. The digital economy is the main economic form of the current society, while R&D funding is the financial support and an important source of green technology innovation. So, these provinces should fully utilize the advantages of the Internet to provide wider and more convenient financing services for technological innovation subjects than ever before. Fourth, remote provinces, such as those belonging to the northeastern region, should take full advantage of the Belt and Road Initiative to strengthen exchanges of talents and international trade cooperation, introduce advanced green technologies and advanced equipment from abroad, and ultimately realize imitation, digestion, and absorption.

Acknowledgments

This research was funded by the Innovation Strategy Research Project of Fujian Province, China, grant number 2023R0075.

Conflict of Interest

The authors declare no conflict of interest.

References

- 1. RAMZAN M., ABBASI K.R., SALMAN A., DAGAR V., ALVARADO R., KAGZI M. Towards the dream of go green: An empirical importance of green innovation and financial depth for environmental neutrality in world's top 10 greenest economies. Technological Forecasting and Social Change, **189**, 122370, **2023**.
- 2. FAROOQ U., WEN J., TABASH M.I., FADOUL M. Environmental regulations and capital investment: Does green innovation allow to grow? International Review of Economics & Finance, **89** (A), 878, **2024**.
- 3. LI K. Can resource endowment and industrial structure drive green innovation efficiency in the context of COP 26? Resources Policy, **82**, 103502, **2023**.
- 4. ZHOU D., LU Z., QIU Y. Do carbon emission trading schemes enhance enterprise green innovation efficiency?

Evidence from China's listed firms. Journal of Cleaner Production, **414**, 137668, **2023**.

- 5. LIU T., YAN W., ZHANG Y. Functional or selective policy? - Research on the relationship between government intervention and enterprise innovation in China. International Review of Economics & Finance, **86**, 82, **2023**.
- 6. LI D., ZENG T. Are China's intensive pollution industries greening? An analysis based on green innovation efficiency. Journal of Cleaner Production, **259**, 120901, **2020**.
- 7. WANG X., LUO G., WANG L. The spatial temporal evolution pattern and influencing factors of green innovation efficiency: Based on provincial panel data of Chinese industrial enterprises. Polish Journal of Environmental Studies, **31** (3), 2317, **2022**.
- 8. ZENG W., LI L., HUANG Y. Industrial collaborative agglomeration, marketization, and green innovation: Evidence from China's provincial panel data. Journal of Cleaner Production, **279**, 123598, **2021**.
- 9. WANG K.L., SUN T.T., XU R.Y., MIAO Z., CHENG Y.H. How does internet development promote urban green innovation efficiency? Evidence from China. Technological Forecasting and Social Change, **184**, 122017, **2022**.
- 10. HUANG H., WANG F., SONG M., BALEZENTIS T., STREIMIKIENE D. Green innovations for sustainable development of China: Analysis based on the nested spatial panel models. Technology in Society, **65**, 101593, **2021**.
- 11. LI G., XUE Q., QIN J. Environmental information disclosure and green technology innovation: Empirical evidence from China. Technological Forecasting and Social Change, **176**, 121453, **2022**.
- 12. VIMALNATH P., TIETZE F., JAIN A., GURTOO A., EPPINGER E., ELSEN M. Intellectual property strategies for green innovations - An analysis of the European Inventor Awards. Journal of Cleaner Production, **377**, 134325, **2022**.
- 13. LIU J., AN K., JANG S.C. Threshold effect and mechanism of tourism industrial agglomeration on green innovation efficiency: Evidence from coastal urban agglomerations in China. Ocean & Coastal Management, **246**, 106908, **2023**.
- 14. ZHAO P., LU Z., KOU J., DU J. Regional differences and convergence of green innovation efficiency in China. Journal of Environmental Management, **325** (A), 116618, **2023**.
- 15. SONG W., HAN X. The bilateral effects of foreign direct investment on green innovation efficiency: Evidence from 30 Chinese provinces. Energy, **261** (B), 125332, **2022**.
- 16. XU Y., LIU S., WANG J. Impact of environmental regulation intensity on green innovation efficiency in the Yellow River Basin, China. Journal of Cleaner Production, **373**, 133789, **2022**.
- 17. CHEN M., SU Y., PIAO Z., ZHU J., YUE X. The green innovation effect of urban energy saving construction: A quasi-natural experiment from new energy demonstration city policy. Journal of Cleaner Production, **428**, 139392, **2023**.
- 18. ZHANG J., KANG L., LI H., BALLESTEROS-PÉREZ P., SKITMORE M., ZUO J. The impact of environmental regulations on urban green innovation efficiency: The case of Xi'an. Sustainable Cities and Society, **57**, 102123, **2020**.
- 19. SONG W., MENG L., ZANG D. Exploring the impact of human capital development and environmental regulations on green innovation efficiency. Environmental Science and Pollution Research, **30**, 67525, **2023**.
- 20. TONE K. A slacks-based measure of efficiency in data envelopment analysis. European Journal of Operational Research, **130** (3), 498, **2001**.
- 21. TONE K. A slacks-based measure of super-efficiency in data envelopment analysis. European Journal of Operational Research, **143** (1), 32, **2002**.
- 22. LI G., LI X., HUO L. Digital economy, spatial spillover and industrial green innovation efficiency: Empirical evidence from China. Heliyon, **9** (1), e12875, **2023**.
- 23. HU Y., WANG C., ZHANG X., WAN H., JIANG D. Financial agglomeration and regional green innovation efficiency from the perspective of spatial spillover. Journal of Innovation & Knowledge, **8** (4), 100434, **2023**.
- 24. LIANG Z., CHEN J., JIANG D., SUN Y. Assessment of the spatial association network of green innovation: Role of energy resources in green recovery. Resources Policy, **79**, 103072, **2022**.
- 25. ZHANG M., HONG Y., WANG P., ZHU B. Impacts of environmental constraint target on green innovation efficiency: Evidence from China. Sustainable Cities and Society, **83**, 103973, **2022**.
- 26. ABUALIGAH L., ELAZIZ M.A., SUMARI P., GEEM Z.W., GANDOMI A.H. Reptile Search Algorithm (RSA): A nature-inspired meta-heuristic optimizer. Expert Systems with Applications, **191**, 116158, **2022**.
- 27. GYEDU S., HENG T., NTARMAH A.H., HE Y., FRIMPPONG E. The impact of innovation on economic growth among G7 and BRICS countries: A GMM style panel vector autoregressive approach. Technological Forecasting and Social Change, **173**, 121169, **2021**.
- 28. ZHOU H., WANG R. Exploring the impact of energy factor prices and environmental regulation on China's green innovation efficiency. Environmental Science and Pollution Research, **29**, 78973, **2022**.
- 29. SUN Y., DING W., YANG G. Green innovation efficiency of China's tourism industry from the perspective of shared inputs: Dynamic evolution and combination improvement paths. Ecological Indicators, **138**, 108824, **2022**.
- 30. WANG K.L., ZHANG F.Q., XU R.Y., MIAO Z., CHENG Y.H., SUN H.P. Spatiotemporal pattern evolution and influencing factors of green innovation efficiency: A China's city level analysis. Ecological Indicators, **146**, 109901, **2023**.
- 31. LI T., SHI Z., HAN D., ZENG J. Agglomeration of the new energy industry and green innovation efficiency: Does the spatial mismatch of R&D resources matter? Journal of Cleaner Production, **383**, 135453, **2023**.
- 32. LEE C.C., NIE C. Place-based policy and green innovation: Evidence from the national pilot zone for ecological conservation in China. Sustainable Cities and Society, **97**, 104730, **2023**.
- 33. DU J., LIANG L., ZHU J. A slacks-based measure of super-efficiency in data envelopment analysis: A comment. European Journal of Operational Research, **204** (3), 694, **2010**.
- 34. DU Q., DENG Y., ZHOU, J., WU J., PANG Q. Spatial spillover effect of carbon emission efficiency in the construction industry of China. Environmental Science and Pollution Research, **29**, 2466, **2022**.
- 35. ALYOUSIFI Y., IBRAHIM K., KANG W., ZIN W.Z.W. Modeling the spatio-temporal dynamics of air pollution index based on spatial Markov chain model. Environmental Monitoring and Assessment, **192**, 719, **2020**.
- 36. WANG Y., CHEN F., WEI F., YANG M., GU X., SUN Q., WANG X. Spatial and temporal characteristics and evolutionary prediction of urban health development efficiency in China: Based on super-efficiency SBM model and spatial Markov chain model. Ecological Indicators, **147**, 109985, **2023**.
- 37. ZHANG Y., LIU Q., LI X., ZHANG X., QIU Z. Spatialtemporal evolution characteristics and critical factors identification of urban resilience under public health emergencies. Sustainable Cities and Society, **102**, 105221, **2024**.
- 38. HUANG B., WU B., BARRY M. Geographically and temporally weighted regression for modeling spatiotemporal variation in house prices. International Journal of Geographical Information Science, **24** (3), 383, **2010**.
- 39. LIU Q., WU R., ZHANG W., LI W., WANG S. The varying driving forces of PM2.5 concentrations in Chinese cities: Insights from a geographically and temporally weighted regression model. Environment International, **145**, 106168, **2020**.
- 40. LI W., JI Z., DONG F. Spatio-temporal evolution relationships between provincial CO_2 emissions and driving factors using geographically and temporally weighted regression model. Sustainable Cities and Society, **81**, 103836, **2022**.
- 41. LI M., HUANG K., XIE X., CHEN Y. Dynamic evolution, regional differences and influencing factors of high-quality development of China's logistics industry. Ecological Indicators, **159**, 111728, **2024**.
- 42. LI M., WANG J. Spatial-temporal distribution characteristics and driving mechanism of green total factor productivity in China's logistics industry. Polish Journal of Environmental Studies, **30** (1), 201, **2021**.
- 43. CHEN Y., LI M., HATAB A.A. A spatiotemporal analysis of comparative advantage in tea production in China. Agricultural Economics - Czech, **66** (12), 550, **2020**.
- 44. XIN X., LYU L., ZHAO Y. Dynamic evolution and trend prediction of multi-scale green innovation in China. Geography and Sustainability, **4** (3), 222, **2023**.